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A new approach for independent component analysis and its application for clustering the economic data

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Abstract: In conventional independent component analysis (ICA) algorithms, the definition of the objective function is typically based on specific dependency criteria. The choice of these criteria significantly influences the performance of the algorithm. This article introduces a general class of dependency criteria, which is based on the cumulative distribution function, to characterise the independence of two variables. Furthermore, an applicable ICA algorithm, grounded in this class and utilising a non-parametric estimator, is proposed. The performance of the proposed algorithm is evaluated and compared with several well-known traditional algorithms, using Amari error estimation calculation as a benchmark. The proposed algorithms have been applied to a real-time series data, serving as a pre-processing clustering method.

Keywords: Amari error; clustering; dependence criteria; independent components analysis.

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1 Introduction

Multivariate data, due to its high dimensionality and complex interdependencies among observed values, presents significant challenges for analysis. Therefore, it becomes essential to segregate the data into different aspects to discern patterns effectively. Independent component analysis (ICA) has been extensively utilised as an unsupervised tool for blind source separation in multivariate statistics, particularly for analysing mixtures of independent source signals (Pfister et al., 2019). ICA was first introduced by Héault in the 1980s, with further developments made by Hyvärinen et al. (2001) and Comon (1994). The application of ICA on signal data has been explored by Comon and Jutten (2010) and Nordhausen and Oja (2018).

ICA is a matrix factorisation technique where the signals captured by each individual matrix factor are optimised to be as mutually independent as possible. This method has been highlighted in various works for its applications in dimensionality reduction, deconvolution, data pre-processing, meta-analysis, and other areas across different data types (Sompairac et al., 2019). The primary objective of ICA is to extract significant components from a dataset, which can encompass various forms such as sounds, stock markets, or images.

ICA has been utilised across multiple fields for the extraction of independent components. These include network analysis for traffic incident detection and estimation (Sheikh and Regan, 2022), optimal portfolio diversification (Lassance et al., 2022), identification of structural autoregressive models in macroeconomics (Moneta and Pallante, 2022), and magnetotelluric sounding data analysis (Zhou et al., 2022). Furthermore, it has been employed for clustering in wireless sensor networks (Shahina and Kumar, 2022), image feature extraction (Shang et al., 2022), and various medical applications, such as brain activity analysis (Tabanfar et al., 2022; Meng et al., 2022; Jayabal et al., 2022), electroencephalogram (EEG) analysis (Lyu and Fu, 2022; Antony et al., 2022), individual molecular characterisation (Rincourt et al., 2022), and the dissociation of biologically single-layer networks (Lipshutz et al., 2022), among others.

ICA has found applications in diverse scientific fields, including medical signals (Xie et al., 2017; Pontifex et al., 2017), audio signals for noise elimination (Du and Swamy, 2014), robust automatic speech discernment model projection (Cho and

Park, 2016), biological assays (Aziz et al., 2016), space image reception (Naik and Kumar, 2011), climate data analysis (Funaro et al., 2001), financial data (Moneta et al., 2010), and socio-economic data, work environment, and employment rate (Tan and Zhang, 2012). It is also used as a pre-processing step in clustering time series data (Rahmanishamsi et al., 2018; Nascimento et al., 2017).

ICA is a computational technique utilised for decomposing a multivariate signal into constituent, additive sub-components. It aims to identify independent components by maximising the statistical independence among the estimated elements. The formulation of the ICA algorithm hinges upon the selection of a metric for independence, with two primary paradigms being the minimisation of mutual information and the maximisation of non-Gaussianity.

This study introduces a novel set of dependency criteria founded on the distribution function to characterise the independence of two random variables and examine their properties. Additionally, an ICA algorithm is proposed based on an approach derived from these novel criteria. To assess the efficacy of the proposed algorithm, a comparative analysis is conducted between the new algorithm and existing methods. Finally, the algorithm is employed as a pre-processing step for ICA in clustering a batch of time series examples.

In the upcoming section, we will delve into the concept of ICA. Following this, an exploration into a dependency criterion based on distribution functions will be introduced and substantiated. Section 4 will unveil several new algorithms designed for ICA. Their comparative performance against existing ICA algorithms will be evaluated in terms of the average of Amari errors via a Monte Carlo simulation survey, presented in Section 5. Section 6 will focus on the application of these algorithms in time series data clustering. Furthermore, this study delves into the application of data clustering methodologies to a set of real-time series data, specifically the gross domestic product (GDP) per capita index, extracted from the statistics of 26 countries spanning from 1975 to 2020. The conclusion of the research is drawn by summarising the results and analysing the structures extracted, as detailed in Section 7.

2 Independent component analysis

Given the complexity and high dimensionality inherent in multivariate data analysis, it often presents challenges in its application. Therefore, it becomes essential to segregate different facets of the data for effective utilisation, particularly in the context of identifying specific patterns in such scenarios (Tharwat, 2021).

In this study, ICA has been employed for signal analysis. Consider each signal as a time-varying vector, denoted as $s_i = (s_{1i}, s_{2i}, \dots, s_{ni})^T$, $i = 1, 2, \dots, d$ represent a signal where n is the number of time steps, s_{ij} is time j of the signal s_i , and d is the number of source signals. Here, (d) also represents independent source signals, and a matrix $S = (s_1, s_2, \dots, s_d)$ is defined, where $S \in \mathbb{R}^{n \times d}$ is matrix of the source signals. These source signals may be mixed, and each source signal influencing the output signals in different ways. Therefore, d mixtures can be represented as $X =: SA$, where $X \in \mathbb{R}^{n \times d}$ is the matrix of the mixed signals, n is the number of mixes, and $A \in \mathbb{R}^{d \times d}$ is a mixing coefficients matrix. S into a mixed signal in $X = (x_1, x_2, \dots, x_d)$ space as $X = SA$. The objective of the model is defined as the extraction of source signals by finding the unmixing coefficient matrix (W). This matrix is used to convert

the mixed signals into a set of independent signals as $X \rightarrow Y : Y = XW$, where $W \in \mathbb{R}^{d \times d}$ is the matrix of unmixing coefficients, and it is inverse of the matrix A .

Based on the estimation principles of ICA, the stages of centralisation and whitening are commonly employed. Data whitening involves transforming mixed signals into uncorrelated signals and scaling them by their respective variances. This whitening process is a step that is essential for preparing the data for the ICA algorithm. Notably, the principal component analysis (PCA) method can be used to whiten the output signals in ICA. The PCA has been used by Ismail et al. (2009), Anbumalar et al. (2011), Kaur and Vashisht (2011), Sivaramakrishna (2011) and Rahmani et al. (2012).

In various research studies, different types of ICA procedures and their interpretations have been explored. Most ICA algorithms aim to minimise a contrast function that measures the degree of dependency between components. The effectiveness of these algorithms is contingent upon the choice of the contrast function and the algorithm used to implement the optimisation problem. The estimation of the unmixing matrix (W matrix) can be achieved through several main approaches to independence, leading to the creation of unmixed matrices with slight variations. In all these approaches, an unmixed matrix is obtained, and the whitened data is projected onto this matrix to extract the independent signals. For further details, the works of Tharwat (2021), Pfister et al. (2019) and Stone (2004) are recommended for reference.

To estimate the independent components, several methods have investigated in previous re- searches. Based on Hyvärinen et al. (2001), maximising non-Gaussianity by negative entropy and estimating the W matrix by the maximum likelihood. Minimising the mutual information (MI) criterion (Hyvärinen, 1999) has represented one of the most popular ICA estimation perspectives. There are several algorithms constructed base on minimising MI to estimate the independent components (Langlois et al., 2010).

Some of the most applicable algorithm in these perspectives are FastICA algorithm based on MI minimisation (Hyvärinen, 1999), Infomax algorithm based on the maximum likelihood method (Lee et al., 1999), JADE algorithm based on maximising kurtosis (Cardoso, 1999), RADICAL algorithm based on the kullback leibler criterion (Learned-Miller and Fisher, 2003), HICA algorithm based on the copula function of Hoeffding's criterion (Rahmanishamsi et al., 2018) and RLICA algorithm based on the copula function of the squared loss MI criterion (Rahmanishamsi and Dolati, 2018). In the next section, some ICA approaches based on minimising MI will be investigated.

The estimation of independent components in ICA involves various methods and criteria. One approach, as proposed by Bingham and Hyvärinen (2000), is based on non-Gaussianity, measured by kurtosis and negative entropy (negentropy), with the aim of maximising non-Gaussianity to find the independent components (Shimizu et al., 2006). Additionally, the maximum likelihood (ML) method is utilised in ICA to estimate the matrix W , providing the best fit for the extracted Y signals – see: Gaeta (1990), Pham et al. (1992) and Pearlmutter and Parra (1996). Due to the mutual information (MI) criterion measures the independence between two random variables, many algorithms for ICA are constructed based on minimising MI to estimate the initial signals (Langlois et al., 2010). Therefore, independent components can be obtained by minimising MI between different components; see Pfister et al. (2019) and Tharwat (2021). In this setting, the matrix W is obtained such that the MI criterion to be close to zero.

3 Proposed dependency criterion

The formulation of the objective function in ICA traditionally relies on density or copula density functions. However, accurately estimating these functions poses significant challenges due to inherent errors in the estimation process. A potential resolution to this issue lies in the utilisation of cumulative distribution functions (CDFs) within ICA methodologies. Unlike non-parametric-based density function estimators, empirical distribution functions (EDFs) offer a notably simpler and faster approach in statistical inference. EDFs exhibit a convergence towards theoretical cumulative distribution functions, enhancing their efficacy in approximating dependencies between variables. The use of empirical distribution functions, which converge to theoretical CDFs, holds promise in evaluating the independence between two random variables. This approach proves particularly valuable when assessing a dependency criterion for characterising independence between variables.

Drawing from these interpretations, an investigation centred on a criterion employing cumulative distribution functions has been undertaken. This criterion aims to characterise the independence between two random variables by evaluating their equality to zero.

To explore this function, consider two random variables, X_1 and X_2 , with a joint distribution function F , and individual marginal distribution functions F_1 and F_2 , respectively. Let $g : \mathbb{R}^+ \rightarrow \mathbb{R}^+$ be a decreasing function for $x < 1$, increasing for $x \geq 1$, and $g(1) = 0$. A generalised dependency criterion denoted as GDC_g , based on cumulative distribution functions, is defined as follows:

$$GDC_g(X_1, X_2) = \sup_{(x_1, x_2) \in \mathbb{R}^2} \left[g \left(\frac{F(x_1, x_2)}{F_1(x_1)F_2(x_2)} \right) \right].$$

Notice that these divergences will vanish if and only if the random variables X_1 and X_2 are independent.

Let $X_{n \times 2}$ be the matrix of observation from random vector (X_1, X_2) with the joint distribution function F and the marginal cumulative distribution functions F_1 and F_2 , respectively. Consider $x_{ij} = [X]_{ij}$, where $i = 1, 2, \dots, n$, $j = 1, 2$. The functions denoted as $\hat{F}_1(x_1) = \frac{\sum_{i=1}^n I(x_{i1} \leq x_1)}{n}$ and $\hat{F}_2(x_2) = \frac{\sum_{i=1}^n I(x_{i2} \leq x_2)}{n}$ correspond to the marginal distribution functions, while $\hat{F}(x_1, x_2) = \frac{\sum_{i=1}^n I(x_{i1} \leq x_1, x_{i2} \leq x_2)}{n}$ represents the joint empirical distribution function. Consequently, the GDC_g can be estimated as:

$$\widehat{GDC}_g(X_1, X_2) = \sup_{(x_1, x_2) \in \mathbb{R}^2} \left[g \left(\frac{\hat{F}(x_1, x_2)}{\hat{F}_1(x_1)\hat{F}_2(x_2)} \right) \right]. \quad (1)$$

4 An ICA algorithm based on \widehat{GDC}_g (GDCICA)

In this section, the estimator denoted as \widehat{GDC}_g for GDC_g in equation (1) has been utilised in the development of several new algorithms for ICA. Let \mathbf{X} be a d -dimensional random vector. ICA aims to find a matrix denoted as $W \in \mathbb{R}^{d \times d}$ such that the components Y_1, Y_2, \dots, Y_d of the new random vector $\mathbf{Y} = \mathbf{X}W$ exhibit minimal statistical dependency. A contrast function within an ICA algorithm is typically based

on a measure of dependency, represented as f , and is defined as $\varphi_X(\cdot)$, satisfying $\varphi_X(W) = f(XW)$. Consequently, the core challenge in ICA involves determining $\widehat{W} \in \mathbb{R}^{d \times d}$ that minimises the contrast function $\varphi_X(W)$.

For statistical independence in ICA problem, pairwise independence refers to a sufficient measure (Comon, 1994). Thus a d -dimensional ICA problem solution defines as a generalised successively solution for a 2-dimensional ICA problems. In other words, a d -dimensional linear transformation described by a $d \times d$ orthogonal matrix R defines as equivalent to a composition of 2-dimensional rotations. Furthermore, the idea for searching a rotation angle such that the corresponding demixed dataset has its GICA_g minimised, is similar to the RADICAL algorithm (Learned-Miller and Fisher, 2003).

Based on previous definitions, the general process of the algorithm for the d -dimensional case is summarised as follows:

Algorithm 4.1:

Input: A matrix $X_{n \times d}$ where rows are mixed signals (centred).

Procedure for calculate GDCICA:

- 1 Whiten the matrix X of the form $Y = X \times Q'$, where Q is a whitening matrix;
- 2 Define the matrices \widehat{W} and \widehat{S} with all entries zero, as the initial matrices for unmixing and source signals, respectively.
- 3 Set $R_{ij}(\theta)$, a rotation matrix of elements (i, j) , of the form

$$R_{ij}(\theta) = \begin{pmatrix} 1 & \cdots & 0 & 0 & \cdots & 0 & 0 & \cdots & 0 \\ \vdots & \ddots & \vdots & \vdots & \ddots & \vdots & \vdots & \ddots & \vdots \\ 0 & \cdots & \cos(\theta) & 0 & \cdots & 0 & -\sin(\theta) & \cdots & 0 \\ 0 & \cdots & 0 & 1 & \cdots & 0 & 0 & \cdots & 0 \\ \vdots & \ddots & \vdots & \vdots & \ddots & \vdots & \vdots & \ddots & \vdots \\ 0 & \cdots & 0 & 0 & \cdots & 1 & 0 & \cdots & 0 \\ 0 & \cdots & \sin(\theta) & 0 & \cdots & 0 & \cos(\theta) & \cdots & 0 \\ \vdots & \ddots & \vdots & \vdots & \ddots & \vdots & \vdots & \ddots & \vdots \\ 0 & \cdots & 0 & 0 & \cdots & 0 & 0 & \cdots & 1 \end{pmatrix}, \quad (2)$$

For all $1 \leq i < j \leq d$, repeat the following steps:

- 4 Define $\widehat{GDC}_g(S_i(\theta), S_j(\theta))$ as a function of θ , where it is assumed that $S^{(i,j)}(\theta)$ which denotes the $(i, j)^{\text{th}}$ columns of $Y \times R_{i,j}(\theta)$, is a sample from $(S_i(\theta), S_j(\theta))$.
- 5 Minimise the function $\widehat{GDC}_g(S_i(\theta), S_j(\theta))$ over $\theta \in [-\frac{\pi}{2}, \frac{\pi}{2}]$ and set

$$\theta_0^{(i,j)} = \operatorname{argmin}_{\theta} \widehat{GDC}_g(S_i(\theta), S_j(\theta)).$$

- 6 Update $(i, j)^{\text{th}}$ columns of \widehat{W} as $R'_{i,j}(\theta_0^{(i,j)}) \times Q$, and $(i, j)^{\text{th}}$ columns of \widehat{S} as $Y \times R(\theta_0^{(i,j)})$.

Output: The unmixing matrix \widehat{W} and the matrix of source signal estimates \widehat{S} .

5 Simulation study

To conduct a simulation study, the Monte Carlo simulation method has employed to measure the performance of GDCICA using various algorithms such as FastICA (Hyvärinen, 1999), Infomax (Lee et al., 1999), JADE (Cardoso, 1999), RADICAL (Learned-Miller and Fisher, 2003), HICA (Rahmanishamsi et al., 2018) and RLICA (Rahmanishamsi and Dolati, 2018) within the R software environment.

The data generation process followed these steps:

- Generate a matrix $S_{n \times d}$, where n denoted the number of observations and d denoted the number of variables ($n \geq d$), also each column comprises a random sample of size n drawn from a specific distribution. Those incorporated 18 distinct one-dimensional densities recommended by Bach and Jordan (2002), encompassing distributions such as student-t, uniform, exponential, a mixture of two Laplace densities, symmetric and non-symmetric Gaussian mixtures. Figure 1 illustrates the density plots of these 18 distributions.
- Generate a random matrix

$$A_{d \times d} = \begin{pmatrix} U_{11} & U_{12} & \cdots & U_{1d} \\ \vdots & \vdots & \vdots & \vdots \\ U_{d1} & U_{d2} & \cdots & U_{dd} \end{pmatrix},$$

where U_{ij} , $i, j = 1, 2, \dots, d$, independently follows a arbitrary random distribution.

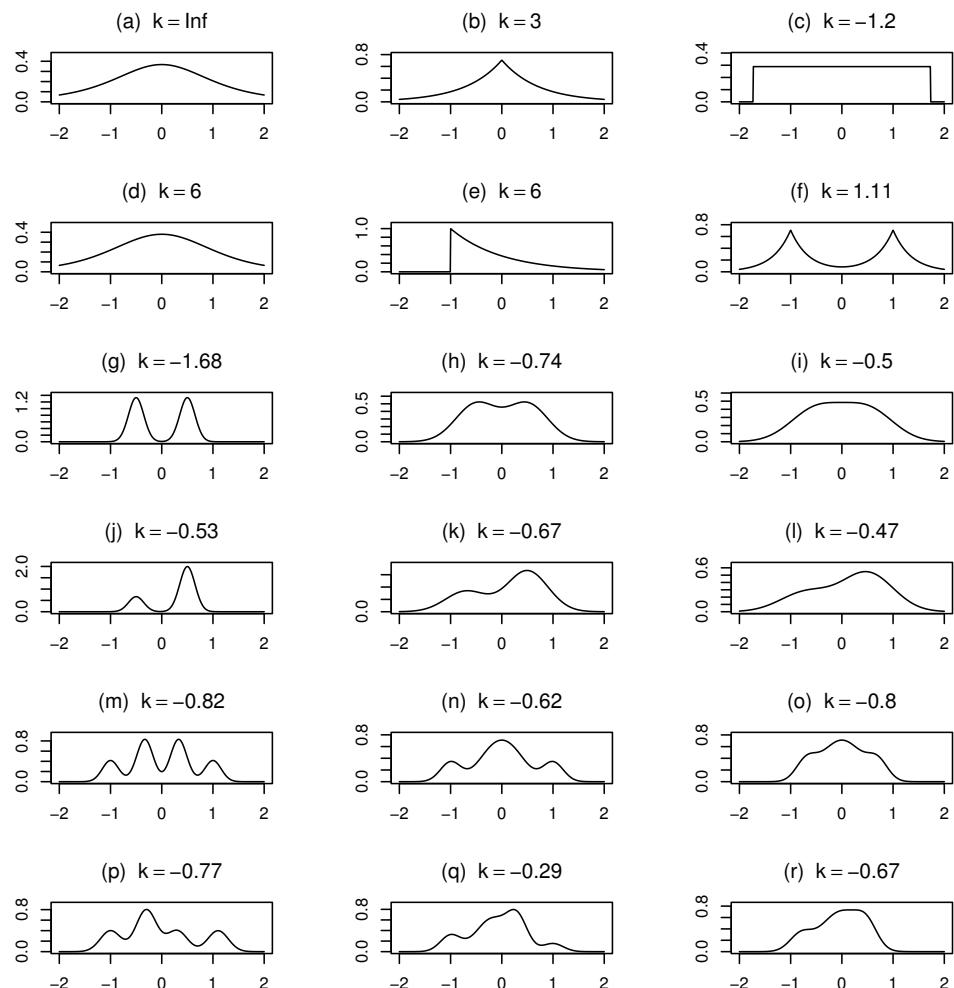
- Construct a data matrix $X_{n \times d}$ using the formula $X = SA$.
- Whiten the matrix X of the form $Y = X \times Q'$. It can be whitened by multiplying X by the inverse of the square root of the sample covariance matrix, resulting in a matrix of whitened data $X_{n \times d}$. This whitened data matrix (Q) is the input for the GDCICA algorithms.

Subsequently, all ICA algorithms generated an unmixing matrix W , which was applied to matrix X to recover estimations of the independent components. The Amari error, initially proposed by Amari et al. (1996), serves as a robust method for error assessment in this context. Let A represent the mixing matrix, while the unmixing matrix is defined through the equation $W = A^{-1}$ using \widehat{W} estimator. The Amari error computation is based on the following expression:

$$\text{Amari error} = \frac{1}{2d(d-1)} \sum_{i,j=1}^d \left(\frac{|a_{ij}|}{\max_i |a_{ij}|} + \frac{|a_{ij}|}{\max_j |a_{ij}|} \right) - \frac{1}{d-1}.$$

Here, the symbol $a_{ij} = [\widehat{W}A]_{ij}$ represents an estimated unmixing matrix \widehat{W} . Under these assumptions, the following conditions hold true within this expression: the metric $[0, d - 1]$ consistently resides within the range of Amari error, also equating to zero solely if \widehat{W} and W denote similar components. Furthermore, this metric remains invariant to the permutation and scaling of columns within A and \widehat{W} .

Figure 1 Density plots of 18 different distribution of sources, (a) student with 3 degrees of freedom (b) double exponential (c) uniform (d) student with 5 degrees of freedom (e) exponential (f) mixture of two double exponentials (g, h, i) symmetric mixtures of two Gaussians: multimodal, transitional and unimodal (j, k, l) non-symmetric mixtures of two Gaussians, multimodal, transitional and unimodal (m, n, o) symmetric mixtures of four Gaussians: multimodal, transitional and unimodal (p, q, r) non-symmetric mixtures of four Gaussians: multimodal, transitional and unimodal



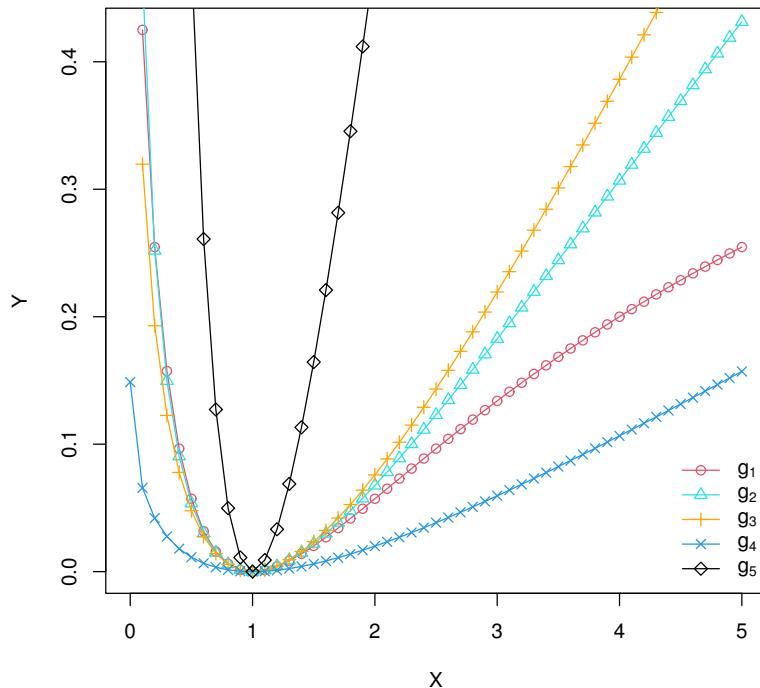
Within GDCICA, certain functions are regarded as special cases of the function g outlined in equations below:

$$\begin{aligned} g_1(x) &= \frac{(x^{c_1} - 1)^2}{x^{2c_1} + 1}, & g_2(x) &= -\ln(x^{c_2}) + x^{c_2} - 1, \\ g_3(x) &= x^{c_3} \ln(x^{c_3}) - x^{c_3} + 1, & g_4(x) &= e^{c_4(x^{c_4^*} - 1)} - c_4(x^{c_4^*} - 1) - 1, \\ g_5(x) &= \ln(x)^{2c_5}, \end{aligned}$$

where $c_1, c_2, c_3, c_4, c_4^*$ and c_5 are some real values.

In GDCICA, a comprehensive exploration of various constants $c_1, c_2, c_3, c_4, c_4^*$ and c_5 has been conducted. The findings reveal that employing specific functions g_1, g_2, g_3, g_4, g_5 with corresponding $c_1 = -0.5, c_2 = 0.5, c_3 = 0.5, c_4 = -0.5, c_4^* = 0.5$, and $c_5 = 1$ values yields commendable performance. Specifically, denoting GDCICA $_g$ with $c_1 = -0.5$ is denoted by GDCICA $_{g_1}$, with $c_2 = 0.5$ is denoted by GDCICA $_{g_2}$, with $c_3 = 0.5$ is denoted by GDCICA $_{g_3}$, with $(c_4, c_4^*) = (-0.5, 0.5)$ is denoted by GDCICA $_{g_4}$, with $c_5 = 1$ is denoted by GDCICA $_{g_5}$. The selected functions are visually presented in Figure 2.

Figure 2 Plots of five different functions in this research (see online version for colours)



In generating 2-dimensional and 4-dimensional datasets of size $n = 250, n = 500, n = 1,000$, and $n = 1,500$, governed by the mixing matrix

$$A_{2 \times 2} = \begin{pmatrix} U_{11} & U_{12} \\ U_{21} & U_{22} \end{pmatrix} \quad \text{and} \quad A_{4 \times 4} = \begin{pmatrix} U_{11} & U_{12} & U_{13} & U_{14} \\ U_{21} & U_{22} & U_{23} & U_{24} \\ U_{31} & U_{32} & U_{33} & U_{34} \\ U_{41} & U_{42} & U_{43} & U_{44} \end{pmatrix},$$

where $U_{ij}, i = 1, 2, 3, 4$ independently follows a uniform distribution over $(0, 2)$, a simulation was conducted. Subsequently, the average Amari errors of methods such as FastICA, Infomax, JADE, RADICAL, HICA, RLICA and GDCICA_g, and others were computed over 120 replications of their respective algorithms.

The outcomes are presented in Tables 1 and 4, illustrating the average errors across various samples. The findings indicate that across 18 cases, each distinct type of GDCICA_g method consistently outperforms its competitors in seven instances, collectively excelling in distributions c, f and j in both 2-dimensional and 4-dimensional situations.

Further elucidation on this investigation is provided below:

- When $n = 250, d = 2$, the GDCICA_{g2}, GDCICA_{g3} and GDCICA_{g4} methods exhibited the lowest Amari errors, that this methods applied to data cases with distributions labelled as a, c, e, f, g, j and m have yielded the best performance, also for $n = 250, d = 4$, the GDCICA_{g2}, GDCICA_{g3} and GDCICA_{g4} methods exhibited the lowest Amari errors, that this methods applied to data cases with distributions labelled as a, c, e, f, j, m and n have yielded the best performance.
- When $n = 500, d = 2$, the GDCICA_{g2} and GDCICA_{g4} demonstrated the most favourable Amari errors, that this methods applied to data cases with distributions labelled as a, c, f, g, j and m have yielded the best performance, also for $n = 500, d = 4$, the GDCICA_{g2} and GDCICA_{g4} demonstrated the most favourable Amari errors, that this methods applied to data cases with distributions labelled as a, c, e, f, j and n have yielded the best performance.
- When $n = 1,000, d = 2$, the GDCICA_{g2} and GDCICA_{g3} showcased the best Amari errors, that this methods applied to data cases with distributions labelled as a, c, e, f, g, j and m have yielded the best performance, also for $n = 1,000, d = 4$, the GDCICA_{g2} and GDCICA_{g3} showcased the best Amari errors, that this methods applied to data cases with distributions labelled as a, c, e, f, g, j, n and r have yielded the best performance.
- When $n = 1,500, d = 2$, the GDCICA_{g2}, GDCICA_{g3} and GDCICA_{g4} again emerged with the most optimal Amari errors, that this methods applied to data cases with distributions labelled as c, f, h, i, j and n have yielded the best performance, also for $n = 1,500, d = 4$, the GDCICA_{g2}, GDCICA_{g3} and GDCICA_{g4} again emerged with the most optimal Amari errors, that this methods applied to data cases with distributions labelled as c, f, i, j, m and n have yielded the best performance.

Further, two distributions were selected from among all distributions labelled (a) to (r). The average Amari errors regarding these selections were computed and discussed in the final row of Tables 1 and 4. The results have revealed that when a stochastic distribution is chosen, the Amari error for the GDCICA_{g2} and GDCICA_{g4} methods yields better performance than other methods.

Table 1 Averages of Amari errors for $d = 2$, $d = 4$, and $n = 250$ sample with 120 replications for each distribution (a) to (r), the smallest (best) entry of our functions in each row is boldfaced

Distribution	d	FastICA	Infomax	JADE	RADICAL	HICA	RICA	GDCICA g_1	GDCICA g_2	GDCICA g_3	GDCICA g_4	GDCICA g_5
a	2	0.5033	0.5072	0.5307	0.5297	0.5052	0.5117	0.5024	0.4895	0.4961	0.5053	0.5034
	4	0.4472	0.4471	0.4525	0.4534	0.4446	0.4328	0.4394	0.4326	0.4328	0.4344	0.4344
b	2	0.4701	0.4705	0.4871	0.4652	0.4740	0.4824	0.4825	0.5084	0.4982	0.4977	0.4854
	4	0.4318	0.4319	0.4822	0.4505	0.4466	0.4535	0.4427	0.4435	0.4365	0.4392	0.4391
c	2	0.5052	0.5049	0.5067	0.4947	0.5042	0.5049	0.4879	0.5044	0.4732	0.4851	0.4790
	4	0.4455	0.4450	0.4456	0.4342	0.4488	0.4522	0.4391	0.4340	0.4340	0.4326	0.4354
d	2	0.4766	0.4886	0.5015	0.5131	0.5059	0.5125	0.5117	0.5345	0.5164	0.5226	0.5198
	4	0.4507	0.4503	0.4644	0.4209	0.4482	0.4526	0.4374	0.4396	0.4356	0.4365	0.4345
e	2	0.5018	0.5019	0.5254	0.5109	0.5208	0.5201	0.5015	0.4851	0.4847	0.4866	0.5017
	4	0.4344	0.4339	0.4458	0.4344	0.4506	0.4516	0.4358	0.4327	0.4332	0.4358	0.4358
f	2	0.5571	0.5570	0.5532	0.5647	0.5611	0.5620	0.5155	0.5094	0.5166	0.5197	0.5150
	4	0.4427	0.4428	0.4426	0.4430	0.4493	0.4475	0.4417	0.4392	0.4362	0.4411	0.4468
g	2	0.5345	0.5347	0.5308	0.5347	0.5408	0.5335	0.5176	0.5123	0.5152	0.5139	0.5150
	4	0.4445	0.4345	0.4451	0.4349	0.4531	0.4537	0.4384	0.4400	0.4352	0.4418	0.4419
h	2	0.5178	0.5313	0.5049	0.5270	0.4977	0.5071	0.5301	0.5330	0.5340	0.5335	0.5292
	4	0.4368	0.4372	0.4621	0.4084	0.4496	0.4512	0.4393	0.4367	0.4388	0.4376	0.4390
i	2	0.5205	0.5161	0.5385	0.5234	0.5238	0.5033	0.5582	0.5439	0.5622	0.5475	0.5549
	4	0.4390	0.4355	0.4617	0.4420	0.4530	0.4527	0.4425	0.4365	0.4402	0.4388	0.4408
j	2	0.4896	0.4893	0.5271	0.4893	0.4808	0.4814	0.4600	0.4603	0.4745	0.4685	0.4701
	4	0.4451	0.4402	0.4475	0.4433	0.4475	0.4456	0.4429	0.4429	0.4370	0.4334	0.4312
k	2	0.5111	0.5107	0.5289	0.5187	0.5305	0.5438	0.5217	0.5095	0.5218	0.5112	0.5358
	4	0.4347	0.4347	0.4474	0.4034	0.4542	0.4500	0.4383	0.4405	0.4371	0.4395	0.4414
l	2	0.4769	0.4925	0.5061	0.5355	0.5466	0.5187	0.5648	0.5527	0.5593	0.5498	0.5568
	4	0.4647	0.4595	0.4433	0.4386	0.4488	0.4580	0.4456	0.4414	0.4399	0.4431	0.4478
m	2	0.4374	0.4372	0.4468	0.4631	0.4642	0.4560	0.4313	0.4238	0.4268	0.4203	0.4321
	4	0.4435	0.4405	0.4344	0.4366	0.4426	0.4492	0.4354	0.4339	0.4327	0.4332	0.4389
n	2	0.4635	0.4577	0.4774	0.4905	0.5285	0.5074	0.5136	0.5156	0.5124	0.5070	0.5104
	4	0.4391	0.4351	0.4325	0.4325	0.4524	0.4532	0.4362	0.4313	0.4307	0.4393	0.4316
o	2	0.5202	0.5231	0.5375	0.5395	0.5328	0.5319	0.5269	0.5260	0.5272	0.5255	0.5255
	4	0.4202	0.4205	0.4313	0.4331	0.4500	0.4501	0.4413	0.4327	0.4347	0.4354	0.4354
p	2	0.5295	0.5165	0.5360	0.4798	0.5046	0.5016	0.4957	0.5141	0.5064	0.4959	0.4916
	4	0.4523	0.4525	0.4501	0.4319	0.4525	0.4529	0.4351	0.4423	0.4371	0.4406	0.4383
q	2	0.5295	0.5165	0.5360	0.4798	0.5046	0.5016	0.4957	0.5141	0.5064	0.4959	0.4916
	4	0.4398	0.4002	0.4332	0.4240	0.4531	0.4448	0.4334	0.4417	0.4391	0.4373	0.4423
r	2	0.5056	0.5010	0.4919	0.5100	0.5061	0.5070	0.5604	0.5588	0.5584	0.5269	0.5596
	4	0.4286	0.4287	0.4315	0.4378	0.4503	0.4512	0.4459	0.4388	0.4372	0.4379	0.4389
rand	2	0.5119	0.5257	0.5175	0.5074	0.5066	0.5152	0.5077	0.4975	0.5014	0.4990	0.5075
	4	0.4314	0.4371	0.4315	0.4394	0.4482	0.4539	0.4369	0.4302	0.4304	0.4307	0.4374

Table 2 Averages of Amari errors for $d = 2$, $d = 4$, and $n = 500$ sample with 120 replications for each distribution (a) to (r), the smallest (best) entry of our functions in each row is boldfaced

Distribution	d	FastICA	Infomax	JADE	RADICAL	HICA	RIICA	GDCICA g_1	GDCICA g_2	GDCICA g_3	GDCICA g_4	GDCICA g_5
a	2	0.5668	0.5668	0.5447	0.5755	0.5538	0.5743	0.5136	0.5232	0.5069	0.5183	0.5191
	4	0.4422	0.4422	0.4756	0.4643	0.4465	0.4494	0.4406	0.4393	0.4450	0.4378	0.4455
b	2	0.4927	0.4925	0.4993	0.4923	0.4698	0.4846	0.4914	0.4977	0.4877	0.4876	0.4880
	4	0.4354	0.4354	0.4543	0.4517	0.4538	0.4596	0.4377	0.4393	0.4438	0.4386	0.4412
c	2	0.4929	0.4929	0.4925	0.5050	0.4978	0.4935	0.4670	0.4646	0.4710	0.4712	0.4692
	4	0.4605	0.4605	0.4850	0.4432	0.4486	0.4505	0.4459	0.4394	0.4385	0.4399	0.4485
d	2	0.5210	0.5244	0.5335	0.5221	0.5599	0.5382	0.5397	0.5329	0.5286	0.5314	0.5374
	4	0.4287	0.4283	0.4254	0.4183	0.4577	0.4503	0.4484	0.4403	0.4404	0.4401	0.4429
e	2	0.5100	0.5105	0.5040	0.5188	0.5305	0.5260	0.5186	0.5448	0.5362	0.5104	0.5354
	4	0.4552	0.4563	0.4642	0.4500	0.4454	0.4507	0.4388	0.4396	0.4419	0.4412	0.4588
f	2	0.5087	0.5086	0.5163	0.5136	0.5227	0.5123	0.4984	0.4879	0.5063	0.5044	0.5031
	4	0.4381	0.4381	0.4487	0.4433	0.4542	0.4521	0.4380	0.4314	0.4378	0.4345	0.4345
g	2	0.5251	0.5250	0.5252	0.5257	0.5133	0.5051	0.5018	0.4968	0.5102	0.5126	0.5126
	4	0.4324	0.4320	0.4324	0.4420	0.4536	0.4543	0.4429	0.4405	0.4448	0.4426	0.4401
h	2	0.5004	0.4835	0.5194	0.5185	0.4854	0.5676	0.5657	0.5771	0.5827	0.5577	0.5577
	4	0.4592	0.4586	0.4646	0.4363	0.4498	0.4522	0.4386	0.4427	0.4487	0.4416	0.4497
i	2	0.4773	0.4713	0.5191	0.5196	0.5341	0.5122	0.5416	0.5293	0.5247	0.5378	0.5443
	4	0.4556	0.4556	0.4578	0.4167	0.4530	0.4460	0.4454	0.4471	0.4422	0.4445	0.4476
j	2	0.5337	0.5072	0.5226	0.5134	0.5106	0.5097	0.4980	0.4916	0.5023	0.5003	0.5045
	4	0.4432	0.4422	0.4432	0.4547	0.4524	0.4438	0.4438	0.4434	0.4330	0.4369	0.4327
k	2	0.5703	0.5629	0.5690	0.5540	0.5517	0.5452	0.5668	0.5717	0.5633	0.5645	0.5557
	4	0.4812	0.4754	0.4811	0.5128	0.4825	0.4874	0.4489	0.4449	0.4387	0.4377	0.4497
l	2	0.4325	0.4324	0.4309	0.4318	0.4459	0.4461	0.4382	0.4335	0.4308	0.4305	0.4328
	4	0.4300	0.4291	0.4348	0.4394	0.4528	0.4518	0.4379	0.4396	0.5086	0.5369	0.5144
m	2	0.4925	0.4923	0.4880	0.4905	0.5526	0.4964	0.4886	0.4885	0.4948	0.4885	0.4860
	4	0.4476	0.4374	0.4324	0.4418	0.4459	0.4561	0.4445	0.4391	0.4418	0.4457	0.4397
n	2	0.4346	0.4222	0.4393	0.4672	0.5105	0.4746	0.5149	0.5110	0.5111	0.5111	0.5014
	4	0.4325	0.4324	0.4309	0.4318	0.4459	0.4561	0.4335	0.4308	0.4305	0.4305	0.4328
o	2	0.5460	0.5461	0.5565	0.5657	0.5672	0.5746	0.5681	0.5679	0.5677	0.5665	0.5771
	4	0.4641	0.4631	0.4507	0.4088	0.4517	0.4500	0.4428	0.4490	0.4410	0.4474	0.4391
p	2	0.5378	0.5368	0.5464	0.5317	0.5109	0.5087	0.5311	0.5273	0.5181	0.5207	0.5336
	4	0.4097	0.4082	0.4350	0.4535	0.4473	0.4445	0.4398	0.4475	0.4492	0.4464	0.4491
q	2	0.5402	0.5409	0.5373	0.5445	0.5394	0.5231	0.5413	0.5560	0.5391	0.5501	0.5570
	4	0.4050	0.4006	0.4358	0.4104	0.4515	0.4538	0.4418	0.4466	0.4409	0.4417	0.4483
r	2	0.4955	0.4955	0.4893	0.4838	0.5061	0.5034	0.5017	0.4997	0.4969	0.4912	0.4905
	4	0.4467	0.4468	0.4529	0.4325	0.4508	0.4552	0.4426	0.4475	0.4403	0.4412	0.4405
rand	2	0.5107	0.5311	0.4944	0.5027	0.4932	0.4968	0.5121	0.4860	0.4942	0.4918	0.4974
	4	0.4332	0.4332	0.4412	0.4382	0.4684	0.4671	0.4454	0.4323	0.4347	0.4325	0.4456

Table 3 Averages of Amari errors for $d = 2$, $d = 4$, and $n = 1,000$ sample with 120 replications for each distribution (a) to (r), the smallest (best) entry of our functions in each row is boldfaced

Distribution	d	FastICA	Infomax	JADE	RADICAL	HICA	RIICA	GDCICA g_1	GDCICA g_2	GDCICA g_3	GDCICA g_4	GDCICA g_5
a	2	0.4878	0.4879	0.4972	0.4941	0.5009	0.5027	0.4737	0.4799	0.4814	0.4735	0.4780
	4	0.4439	0.4439	0.4383	0.4354	0.4476	0.4441	0.4475	0.4323	0.4305	0.4365	0.4511
b	2	0.5090	0.5090	0.5038	0.5216	0.5066	0.5223	0.5575	0.5347	0.5485	0.5449	0.5374
	4	0.4400	0.4400	0.4178	0.4371	0.4515	0.4527	0.4539	0.4414	0.4366	0.4301	0.4548
c	2	0.5195	0.5195	0.5116	0.5151	0.5128	0.5053	0.4986	0.4851	0.4938	0.4962	0.4938
	4	0.4482	0.4482	0.4461	0.4396	0.4490	0.4483	0.4496	0.4389	0.4367	0.4302	0.4471
d	2	0.4639	0.4639	0.4655	0.4674	0.4656	0.4671	0.4699	0.4664	0.4726	0.4783	0.4897
	4	0.4558	0.4558	0.4442	0.4344	0.4534	0.4547	0.4475	0.4423	0.4430	0.4460	0.4488
e	2	0.5325	0.5446	0.5482	0.5476	0.5418	0.5408	0.5370	0.5240	0.5338	0.5355	0.5404
	4	0.4592	0.4592	0.4566	0.4527	0.4482	0.4492	0.4414	0.4570	0.4449	0.4447	0.4447
f	2	0.5261	0.5260	0.5268	0.5281	0.5280	0.5265	0.5535	0.5602	0.5614	0.5200	0.5679
	4	0.4556	0.4556	0.4407	0.4373	0.4493	0.4554	0.4502	0.4367	0.4395	0.4478	0.4475
g	2	0.5577	0.5546	0.5508	0.5614	0.5614	0.5619	0.5495	0.5555	0.5390	0.5520	0.5492
	4	0.5260	0.5224	0.5283	0.5063	0.4562	0.4504	0.4500	0.4432	0.4455	0.4551	0.4494
h	2	0.4639	0.4639	0.4616	0.4550	0.4594	0.4707	0.4770	0.4823	0.4771	0.4798	0.4704
	4	0.4384	0.4384	0.4270	0.4327	0.4536	0.4478	0.4523	0.4438	0.4399	0.4516	0.4462
i	2	0.4936	0.5028	0.5013	0.4981	0.5145	0.5313	0.5308	0.5246	0.5281	0.5069	0.5352
	4	0.5019	0.4986	0.4839	0.4321	0.4497	0.4491	0.4503	0.4372	0.4427	0.4510	0.4475
j	2	0.6040	0.5740	0.5317	0.5262	0.5279	0.5255	0.5319	0.5446	0.5195	0.5108	0.5245
	4	0.4634	0.4633	0.4482	0.4498	0.4607	0.4533	0.4495	0.4443	0.4380	0.4485	0.4488
k	2	0.5062	0.5062	0.5082	0.5127	0.4984	0.4952	0.5280	0.5336	0.5255	0.5263	0.5153
	4	0.4217	0.4248	0.4757	0.4269	0.4424	0.4404	0.4436	0.4342	0.4346	0.4349	0.4517
l	2	0.5033	0.5039	0.5169	0.5303	0.5387	0.5254	0.5525	0.5509	0.5529	0.5525	0.5455
	4	0.5096	0.5090	0.4950	0.4739	0.4844	0.4723	0.4951	0.4767	0.4975	0.4880	0.4476
m	2	0.4803	0.4803	0.4609	0.4292	0.4468	0.4558	0.4482	0.4457	0.4385	0.4342	0.4488
	4	0.5231	0.5231	0.5019	0.4728	0.5292	0.4692	0.5085	0.4978	0.5025	0.4988	0.5072
n	2	0.4821	0.4841	0.4441	0.4425	0.4465	0.4500	0.4499	0.4368	0.4424	0.4562	0.4523
	4	0.4892	0.4892	0.4877	0.4858	0.5024	0.4714	0.4812	0.4744	0.4940	0.4845	0.4948
o	2	0.4378	0.4325	0.4363	0.4479	0.4479	0.4489	0.4538	0.4488	0.4386	0.4513	0.4557
	4	0.4982	0.4982	0.4923	0.4993	0.5194	0.4931	0.5330	0.5292	0.5255	0.5239	0.5278
p	2	0.4747	0.4742	0.4304	0.4327	0.4523	0.4547	0.4522	0.4393	0.4391	0.4472	0.4481
	4	0.5603	0.5450	0.5197	0.4844	0.4983	0.5190	0.5627	0.5642	0.5419	0.5553	0.5532
q	2	0.4337	0.4337	0.4139	0.4328	0.4493	0.4558	0.4506	0.4423	0.4389	0.4515	0.4489
	4	0.4978	0.4978	0.4741	0.4868	0.4918	0.5075	0.4949	0.5098	0.5118	0.4932	0.4462
r	2	0.4694	0.4694	0.4447	0.4352	0.4491	0.4527	0.4502	0.4371	0.4369	0.4529	0.4462
	4	0.5091	0.5185	0.5290	0.5354	0.5170	0.5351	0.5147	0.5061	0.5058	0.5163	0.5104
rand	2	0.4494	0.4522	0.4329	0.4328	0.4643	0.4616	0.4463	0.4325	0.4334	0.4425	0.4408

Table 4 Averages of Amari errors for $d = 2$, $d = 4$, and $n = 1,500$ sample with 120 replications for each distribution (a) to (r), the smallest (best) entry of our functions in each row is boldfaced

Distribution	d	FastICA	Infomax	JADE	RADICAL	HICA	RIICA	GDCICA g_1	GDCICA g_2	GDCICA g_3	GDCICA g_4	GDCICA g_5
a	2	0.5204	0.5203	0.5266	0.5155	0.5285	0.5242	0.5586	0.5472	0.5488	0.5487	0.5376
	4	0.4436	0.4437	0.4491	0.4552	0.4530	0.4555	0.4454	0.4489	0.4432	0.4455	0.4465
b	2	0.5606	0.5606	0.5708	0.5533	0.5689	0.5590	0.5413	0.5428	0.5359	0.5487	0.5297
	4	0.4405	0.4436	0.4352	0.4452	0.4458	0.4456	0.4408	0.4420	0.4430	0.4393	0.4439
c	2	0.5247	0.5247	0.5244	0.5018	0.5231	0.5258	0.5212	0.5327	0.5195	0.5163	0.5204
	4	0.4438	0.4437	0.4519	0.4524	0.4532	0.4526	0.4482	0.4412	0.4404	0.4405	0.4396
d	2	0.5492	0.5491	0.5554	0.5674	0.5614	0.5606	0.5557	0.5599	0.5570	0.5676	0.5495
	4	0.4444	0.4569	0.4538	0.4482	0.4589	0.4571	0.4515	0.4501	0.4519	0.4520	0.4540
e	2	0.5025	0.5025	0.5020	0.5090	0.5092	0.5050	0.5050	0.5174	0.5012	0.5126	0.5126
	4	0.4423	0.4506	0.4459	0.4525	0.4563	0.4578	0.4533	0.4517	0.4547	0.4551	0.4498
f	2	0.5004	0.5004	0.5001	0.4964	0.5008	0.5017	0.5037	0.5097	0.5060	0.5208	0.5140
	4	0.4528	0.4534	0.4489	0.4554	0.4536	0.4587	0.4447	0.4473	0.4452	0.4497	0.4488
g	2	0.5678	0.5682	0.5662	0.5662	0.5662	0.5632	0.5373	0.5523	0.5475	0.5686	0.5394
	4	0.4345	0.4344	0.4534	0.4589	0.4559	0.4456	0.4570	0.4467	0.4502	0.4463	0.4511
h	2	0.5424	0.5424	0.5415	0.5380	0.5368	0.5279	0.5274	0.5225	0.5092	0.5254	0.5436
	4	0.5396	0.4284	0.4547	0.4548	0.4502	0.4491	0.4494	0.4469	0.4487	0.4507	0.4491
i	2	0.5056	0.5055	0.5134	0.4889	0.4891	0.5041	0.4824	0.4820	0.4879	0.4939	0.4758
	4	0.4470	0.4496	0.4563	0.4560	0.4560	0.4557	0.4485	0.4459	0.4440	0.4453	0.4435
j	2	0.5594	0.5524	0.5459	0.5457	0.5492	0.5454	0.5497	0.5455	0.5522	0.5387	0.5536
	4	0.4539	0.4538	0.4513	0.4566	0.4568	0.4480	0.4496	0.4463	0.4476	0.4437	0.4417
k	2	0.5080	0.5081	0.5026	0.5129	0.5039	0.5127	0.5344	0.5345	0.5309	0.5181	0.5351
	4	0.4599	0.4587	0.4551	0.4589	0.4523	0.4451	0.4519	0.4497	0.4481	0.4526	0.4470
l	2	0.5337	0.5338	0.5268	0.5216	0.5426	0.5404	0.5312	0.5411	0.5375	0.5394	0.5347
	4	0.4451	0.4348	0.4572	0.4514	0.4527	0.4525	0.4525	0.4505	0.4539	0.4572	0.4506
m	2	0.5272	0.5273	0.5143	0.5189	0.5430	0.5181	0.5220	0.5171	0.5139	0.5236	0.5139
	4	0.4440	0.4439	0.4543	0.4570	0.4558	0.4526	0.4549	0.4419	0.4435	0.4425	0.4451
n	2	0.5060	0.5066	0.5193	0.4954	0.5404	0.5200	0.5257	0.5265	0.5306	0.5176	0.5233
	4	0.4451	0.4427	0.4536	0.4517	0.4498	0.4545	0.4504	0.4407	0.4426	0.4384	0.4359
o	2	0.4441	0.4440	0.4408	0.4561	0.4409	0.4358	0.4629	0.4507	0.4495	0.4508	0.4623
	4	0.4540	0.4549	0.4468	0.4523	0.4600	0.4539	0.4515	0.4468	0.4531	0.4544	0.4522
p	2	0.4999	0.4999	0.5069	0.5257	0.5088	0.5061	0.5059	0.5034	0.5062	0.5047	0.5047
	4	0.4403	0.4492	0.4598	0.4542	0.4535	0.4589	0.4435	0.4404	0.4486	0.4434	0.4560
q	2	0.4880	0.4881	0.4826	0.4929	0.4839	0.4927	0.5144	0.5145	0.5109	0.4981	0.5151
	4	0.4439	0.4436	0.4534	0.4553	0.4512	0.4548	0.4532	0.4566	0.4573	0.4489	0.4533
r	2	0.5360	0.5352	0.5545	0.5510	0.5557	0.5524	0.5355	0.5058	0.5335	0.5380	0.5201
	4	0.4389	0.4348	0.4538	0.4577	0.4554	0.4488	0.4475	0.4479	0.4463	0.4514	0.4518
rand	2	0.5360	0.5352	0.4440	0.4556	0.4468	0.4439	0.4437	0.4425	0.4429	0.5201	0.5380
	4	0.4455	0.4435								0.4444	0.4444

Table 5 Averages of Amari errors for $d = 2$, $d = 4$ and $n = 250$, $n = 500$, $n = 1,000$, $n = 1,500$ samples in each class of distributions, the smallest entry of our functions in each row is boldfaced

Class	d	n	FastICA	Infomax	JADE	RADICAL	HICA	RUICA	GDCICA $_{g_1}$	GDCICA $_{g_2}$	GDCICA $_{g_3}$	GDCICA $_{g_4}$	GDCICA $_{g_5}$
symmetric	2	250	0.5006	0.5028	0.5067	0.5128	0.5154	0.5097	0.5084	0.5104	0.5064	0.5065	
		500	0.5053	0.5039	0.5089	0.5178	0.5284	0.5145	0.5178	0.5155	0.5161	0.5188	0.5183
		1000	0.5013	0.5021	0.4987	0.4976	0.5101	0.5001	0.5087	0.5061	0.5087	0.5021	0.5086
	4	1500	0.5226	0.5226	0.5250	0.5180	0.5272	0.5219	0.5217	0.5221	0.5191	0.5248	0.5199
		250	0.4401	0.4382	0.4506	0.4334	0.4498	0.4510	0.4388	0.4370	0.4352	0.4368	0.4381
		500	0.4451	0.4440	0.4507	0.4362	0.4510	0.4524	0.4417	0.4399	0.4413	0.4415	0.4420
non-symmetric	1000	1000	0.4646	0.4642	0.4484	0.4437	0.4510	0.4500	0.4500	0.4400	0.4439	0.4457	0.4499
		1500	0.4536	0.4545	0.4508	0.4534	0.4538	0.4528	0.4493	0.4453	0.4460	0.4459	0.4463
		250	0.5055	0.5039	0.5204	0.5109	0.5106	0.5051	0.5031	0.5009	0.5150	0.5018	0.5119
	1500	500	0.5166	0.5111	0.5135	0.5175	0.5133	0.5097	0.5165	0.5226	0.5171	0.5192	0.5260
		1000	0.5327	0.5273	0.5213	0.5103	0.5188	0.5159	0.5374	0.5346	0.5306	0.5316	0.5308
		1500	0.5199	0.5187	0.5198	0.5243	0.5252	0.5240	0.5177	0.5213	0.5234	0.5186	0.5235
	4	250	0.4428	0.4436	0.4427	0.4432	0.4510	0.4506	0.4395	0.4392	0.4437	0.4379	0.4393
		500	0.4439	0.4438	0.4500	0.4436	0.4502	0.4510	0.4414	0.4420	0.4481	0.4423	0.4461
		1000	0.4617	0.4619	0.4521	0.4431	0.4510	0.4511	0.4430	0.4401	0.4438	0.4487	0.4473
		1500	0.4543	0.4544	0.4537	0.4552	0.4540	0.4502	0.4490	0.4509	0.4509	0.4500	0.4503

Table 5 Averages of Amari errors for $d = 2$, $d = 4$ and $n = 250$, $n = 500$, $n = 1,000$, $n = 1,500$ samples in each class of distributions, the smallest entry of our functions in each row is boldfaced (continued)

Class	d	n	F_{SisilCA}	$Infomax$	$JADE$	$RADICAL$	$HICA$	$RLCA$	$GDCIC_{g_l}$	$GDCIC_{g_2}$	$GDCIC_{g_3}$	$GDCIC_{g_4}$	$GDCIC_{g_5}$
unimodal	2	250	0.4969	0.5005	0.5092	0.5158	0.5183	0.5103	0.5278	0.5241	0.5243	0.5171	0.5261
		500	0.5115	0.5103	0.5159	0.5238	0.5255	0.5251	0.5228	0.5264	0.5194	0.5232	0.5259
		1000	0.4971	0.4999	0.5021	0.5024	0.5072	0.5066	0.5138	0.5100	0.5151	0.5110	0.5143
	4	1500	0.5190	0.5189	0.5238	0.5204	0.5248	0.5230	0.5210	0.5184	0.5189	0.5223	0.5153
		250	0.4396	0.4384	0.4516	0.4375	0.4502	0.4518	0.4405	0.4380	0.4362	0.4399	0.4387
		500	0.4447	0.4449	0.4520	0.4352	0.4515	0.4516	0.4418	0.4427	0.4430	0.4424	0.4453
multimodal	2	1000	0.4672	0.4667	0.4516	0.4316	0.4503	0.4503	0.4507	0.4403	0.4380	0.4456	0.4493
		1500	0.4445	0.4461	0.4499	0.4523	0.4547	0.4533	0.4489	0.4480	0.4488	0.4498	0.4490
		250	0.5054	0.5042	0.5108	0.5021	0.5131	0.5130	0.5098	0.5153	0.5096	0.5065	0.5092
	4	500	0.5077	0.5039	0.5043	0.5180	0.5236	0.5044	0.5315	0.5338	0.5340	0.5359	0.5282
		1000	0.5146	0.5115	0.5006	0.4880	0.4996	0.4919	0.5150	0.5126	0.5082	0.5113	0.5080
		1500	0.5138	0.5140	0.5141	0.5082	0.5176	0.5158	0.5246	0.5261	0.5202	0.5151	0.5275
transitional	2	250	0.4456	0.4421	0.4444	0.4375	0.4490	0.4498	0.4387	0.4384	0.4396	0.4380	0.4394
		500	0.4342	0.4316	0.4381	0.4448	0.4511	0.4519	0.4418	0.4439	0.4413	0.4438	0.4439
		1000	0.4561	0.4553	0.4354	0.4346	0.4349	0.4324	0.4495	0.4404	0.4439	0.4491	0.4479
	4	1500	0.4451	0.4470	0.4555	0.4564	0.4551	0.4527	0.4499	0.4455	0.4470	0.4453	0.4485
		250	0.5096	0.5051	0.5196	0.5055	0.5103	0.5069	0.4840	0.4879	0.4837	0.4848	
		500	0.5196	0.5140	0.5197	0.5149	0.5245	0.5081	0.5042	0.4994	0.5037	0.5131	
	2	1000	0.5333	0.5272	0.5203	0.5178	0.5333	0.5159	0.5326	0.5332	0.5286	0.5189	0.5281
		1500	0.5309	0.5296	0.5271	0.5306	0.5336	0.5269	0.5237	0.5256	0.5258	0.5296	0.5271
		250	0.4439	0.4439	0.4442	0.4392	0.4516	0.4503	0.4372	0.4368	0.4359	0.4372	0.4379
	4	500	0.4477	0.4463	0.4603	0.4393	0.4489	0.4523	0.4409	0.4396	0.4390	0.4388	0.4453
		1000	0.4448	0.4458	0.4414	0.4349	0.4482	0.4548	0.4492	0.4392	0.4343	0.4347	0.4492
		1500	0.4562	0.4544	0.4537	0.4546	0.4513	0.4512	0.4506	0.4470	0.4474	0.4462	0.4450

Table 5 Averages of Amari errors for $d = 2$, $d = 4$ and $n = 250$, $n = 500$, $n = 1,000$, $n = 1,500$ samples in each class of distributions, the smallest entry of our functions in each row is boldfaced (continued)

Class	d	n	$E_{\text{est}}[\text{ICA}]$	$Infomax$	$JADE$	$RADICAL$	$HICA$	$RICA$	$GDC[\mathcal{A}_g]$	$GDC[\mathcal{C}\mathcal{A}_g]$	$GDC[\mathcal{C}\mathcal{A}_{g_i}]$	$GDC[\mathcal{C}\mathcal{A}_{g_3}]$	$GDC[\mathcal{C}\mathcal{A}_{g_5}]$
kurtoses +	2	250	0.5018	0.5050	0.5134	0.5169	0.5183	0.5164	0.5009	0.5046	0.5080	0.5011	0.5054
		500	0.5198	0.5206	0.5196	0.5245	0.5273	0.5271	0.5104	0.5123	0.5173	0.5124	0.5166
		1000	0.5039	0.5063	0.5083	0.5118	0.5086	0.5119	0.5104	0.5183	0.5190	0.5195	0.5227
	4	1500	0.5299	0.5301	0.5310	0.5283	0.5338	0.5305	0.5354	0.5329	0.5354	0.8279	0.5281
		250	0.4413	0.4412	0.4412	0.4575	0.4438	0.4497	0.4500	0.4380	0.4388	0.4348	0.4366
		500	0.4440	0.4401	0.4336	0.4455	0.4515	0.4524	0.4407	0.4380	0.4354	0.4410	0.4446
kurtoses -	2	1000	0.4509	0.4509	0.4440	0.4436	0.4509	0.4510	0.4496	0.4388	0.4373	0.4411	0.4490
		1500	0.4480	0.4496	0.4477	0.4513	0.4535	0.4549	0.4471	0.4480	0.4476	0.4472	0.4486
		250	0.5032	0.5019	0.5122	0.5062	0.5132	0.5076	0.5064	0.5130	0.5115	0.5136	0.5117
	4	500	0.5099	0.5053	0.5115	0.5179	0.5237	0.5098	0.5258	0.5231	0.5224	0.5217	0.5254
		1000	0.5152	0.5124	0.5062	0.4992	0.5138	0.5025	0.5144	0.5197	0.5164	0.5159	0.5140
		1500	0.5187	0.5181	0.5186	0.5165	0.5219	0.5188	0.5173	0.5188	0.5176	0.5179	0.5194

Certainly, the average Amari errors were computed for performance evaluation across all methods in various distributions, categorised into different classes. Initially, the classifications were divided into unimodal distributions (a, b, d, e, i, l, o, r), multimodal distributions (f, g, j, m, p), and transitional distributions (c, h, k, n, q) for the first analysis. Subsequently, the perspective shifted to symmetric distributions (a, b, c, d, f, g, h, i, m, n, o) and non-symmetric distributions (e, j, k, l, p, q, r). Finally, the third perspective involved positively kurtosis distributions (a, b, d, e, f) and negatively kurtosis distributions (c, g, h, i, j, k, l, m, n, o, p, q, r). Each of these classes formed a partition within the set comprising all distributions. The detailed results of this analysis are presented in Table 5.

According to Table 5, the optimal performance in terms of Amari errors was observed in two specific scenarios:

- For $d = 2$, the GDCICA_{g_3} and GDCICA_{g_5} methods have demonstrated superior results in the transitional distributions and positively kurtosis class. Also for $d = 4$, the GDCICA_{g_3} method has demonstrated superior results in the transitional distributions and positively kurtosis class.
- For $d = 2, 4$, the GDCICA_{g_1} , GDCICA_{g_2} and GDCICA_{g_4} methods have demonstrated superior results in the non-symmetric distributions, transitional distributions and positively kurtosis class.

Overall, there was a consistent trend of achieving commendable results in terms of Amari errors across transitional distributions and the positively kurtosis class, indicating a notably favourable performance level.

6 Real data

The application of ICA within statistics involves transforming raw time series data into a set of independent variable values. Its utility extends to cluster analysis, aiding in the identification of datasets exhibiting similar temporal patterns. Moreover, ICA serves as an effective data pre-processing step, enhancing temporal data clustering.

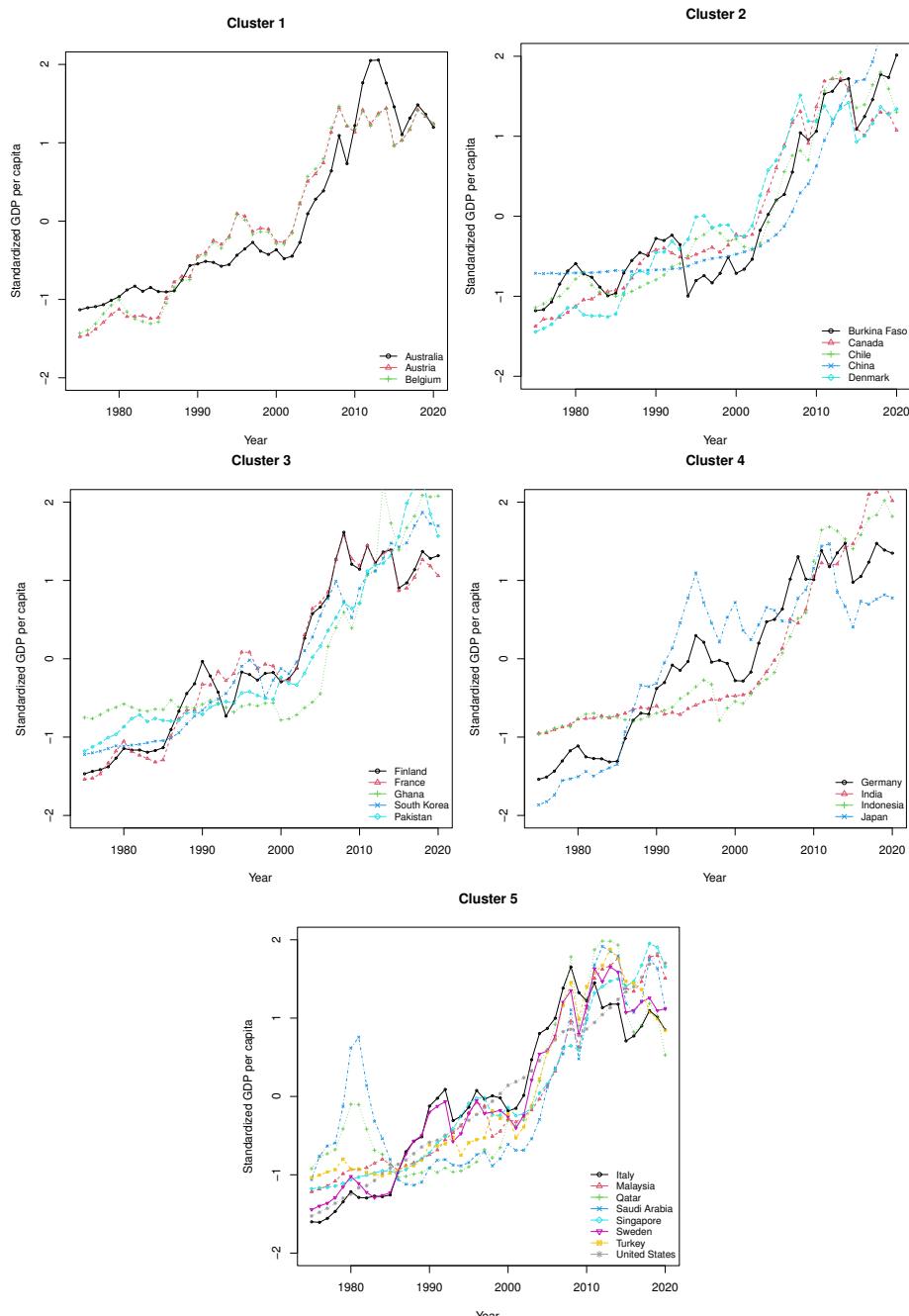
ICA's significance lies in its ability to facilitate clustering for small, distribution-free data series while considering dependencies between successive time points (Nascimento et al., 2017). By inherently addressing temporal dependencies during variable decomposition, ICA swiftly generates independent components. These components are subsequently grouped based on the temporal relationships among them (Calhoun and Adali, 2012). This inherent capacity to account for temporal dependency within its model distinguishes ICA's effectiveness in swiftly generating and clustering independent components, bolstering its relevance in diverse statistical applications.

Based on Guo et al. (2008), Zanghaei et al. (2013) and Rahmanishamsi et al. (2018), instead of applying clustering algorithms on data, it is better that a suitable transformation being used before clustering process to strengthen clustering results. In this article, the ICA algorithm was first implemented on the data and then clustering was done.

Absolutely, GDCICA algorithms offer substantial value in the realm of time series data pre-processing. Their efficacy lies in identifying homogeneous structures within the data, thereby maximising within-group similarities while emphasising distinctions

between different groups. This approach enhances the effectiveness of time series clustering, rendering it an attractive and potent method for extracting valuable insights across diverse domains.

Figure 3 Trend plots of standardised GDP per capita time series in five clusters (see online version for colours)



ICA's applications extend beyond time series clustering; it provides valuable insights into common characteristics of datasets. This method involves a multi-stage process utilising the GDCICA algorithm as a pre-processing step in time series clustering when applied to real data. Statistically independent components can be obtained through GDCICA, leveraging dependency implications to reduce data dimensionality. This method effectively detects dependencies within time series data. By generating new components from sources containing ample information about time series trends, this pre-processing technique enhances data representation. Consequently, the coefficients of the mixing matrix derived from GDCICA can serve as input variables in clustering methods.

Let X be a time series matrix and \hat{S} and \hat{A} be estimates of the source signal matrix and the mixing coefficients matrix obtained by an ICA algorithm on X , respectively. Thus, time series matrix X can be predicted by $\hat{S}\hat{A}$; i.e., $X \approx \hat{S}\hat{A}$, which it means that matrix X is approximately equal to $\hat{S}\hat{A}$. We know that the columns of \hat{S} including the independent time series and the their dependencies are converted to the weight matrix \hat{A} . On the other hand, i -th column of X is constructed as multiplication of matrix \hat{S} and the i -th column of \hat{A} . Thus, the same columns in \hat{A} construct the same trends in the original time series. Therefore, detecting the similar columns of \hat{A} , imply to detect the similar columns of X . Hence, for clustering the time series matrix X , we can apply a clustering algorithm on \hat{A} .

Table 6 Clustering of standardised GDP per capita time series.

Method	Suggested number of clusters	Average silhouette	Method	Suggested number of clusters	Average silhouette
GDCICA _{g1}	5	0.092	JADE	7	0.069
GDCICA _{g2}	3	0.118	Infomax	6	0.049
GDCICA _{g3}	5	0.129	FastICA	8	0.059
GDCICA _{g4}	3	0.104	RADICAL	6	0.055
GDCICA _{g5}	3	0.110	HICA	7	0.080
			RLICA	4	0.090

Then, the GDP per capita time series data for 25 countries: Australia, Austria, Belgium, Burkina Faso, Canada, Chile, China, Denmark, Finland, France, Germany, Ghana, India, Indonesia, Italy, Japan, South Korea, Malaysia, Pakistan, Qatar, Saudi Arabia, Singapore, Sweden, Turkey, and the USA, was sourced from the World Bank's repository (<https://data.worldbank.org>) spanning the period 1975 to 2020.

In this study, time series clustering employed suggested algorithms as a pre-processing step. Subsequently, the PAM algorithm (k -medoids), introduced by Rdušeeun and Kaufman (1987), was applied for final clustering.

Furthermore, after estimating the mixing matrix using multiple algorithms, including GDCICA with FastICA, Infomax, JADE, RADICAL, HICA and RLICA, the PAM clustering algorithm was employed on the unmixing matrix. This step aimed to determine the appropriate number of clusters and declare the best clustering method based on the silhouette criterion, utilising the NbClust R package (Rdušeeun and Kaufman, 1987). The detailed results are reported in Table 6. Based on the results, the function exhibiting the highest average silhouette score has been selected as the best-performing result.

This dataset was subjected to clustering using the GDCICA_{g_3} function. The primary objective of this experiment was to cluster countries based on the patterns observed in their GDP per capita time series spanning the past 46 years. To the study objective satisfaction, the primary GDP per capita data has standardised. Then, the GDCICA algorithm has applied on the principal components and the coefficients of the mixing matrix in GDP per capita. The extracted data has utilised as input to PAM clustering algorithm. The clustering results illustrated in Table 7.

Table 7 Clustering of standardised GDP per capita time series.

Cluster 1	Australia, Austria, Belgium.
Cluster 2	Burkina Faso, Canada, Chile, China, Denmark.
Cluster 3	Finland, France, Ghana, South Korea, Pakistan.
Cluster 4	Germany, India, Indonesia, Japan.
Cluster 5	Italy, Malaysia, Qatar, Saudi Arabia, Singapore, Sweden, Turkey, USA.

Although these time series data adhere to various distributions, the proposed generalised density contrast independent component analysis (GDCICA) algorithms are capable of effectively extracting the sources necessary for clustering. To evaluate the method's efficiency, time series plots within five distinct clusters have been generated, as depicted in Figure 3. The trend plots derived from all clusters, based on the standardised gross domestic product (GDP) per capita, indicate that the preprocessing technique employed by the proposed algorithm for clustering countries is highly effective. This is evidenced by the clear observation that countries within the same cluster exhibit similar trends in their standardised GDP per capita.

7 Conclusions

The objective function of many common ICA algorithms is defined based on density functions, which estimate marginal and joint density functions. This can lead to longer and time-consuming algorithm execution. As a result, this is a challenging problem. In this study, a class of dependence criteria of two continues random variables X_1 and X_2 based on the cumulative distribution function, called GDC has proposed to solve the challenging problem. The criterion of this class remains unchanged despite increasing transformations of random variables, and it only disappears when the random variables X_1 and X_2 are independent.

Then, some algorithms for ICA problem based on this dependence criterion has presented called GDCICA. To suggested algorithms performance evaluation, these has compared to some exiting algorithms using by the Monte-Carlo simulation studies, when the source signals follow the 18 different distributions. The simulation results argued that the suggested algorithms estimate the unmixing matrix very well. So, it obtained better performance than the usual ICA algorithms in various classes of distributions bas on average of the Amari errors. Usually, various types of GDCICA_g have good performance in term of the average of the Amari errors in the transitional distributions and positively kurtosis class.

So, one batch of GDP per capita for 25 countries during 1975 to 2020 has utilised as time series samples to clustering. At first, all data standardised. Then, the GDCICA has

utilised on principal components as a pre-processing stage. As a result of this stage, the coefficients of the matrix have extracted. Then, they were applied as input to the PAM clustering algorithm. The clustering results concluded from a time series data argued which the pre-processing technique using proposed algorithm are fantastic in results of suitable clustering of different data that follow different distributions.

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